

Poisson Mixture Models

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Much of this material is adapted from Bilmes 1998 and Tomasi 2004.

Many of the images were taken from the Internet

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Poisson Mixture Models

Suppose we have a dataset \mathcal{D} which consists of DNA sequences observed from a mixture of k bacteria. We do not know which sequence belongs to which species.

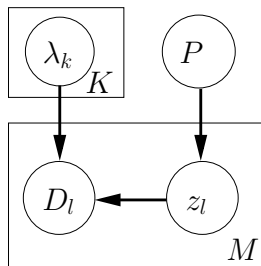
| Sequence | Species | Count |
|----------|---------|-------|
| CAGAGGAT | ? | 5 |
| TCAGTGTC | ? | 13 |
| CTCTGTGA | ? | 2 |
| AACTGTCG | ? | 7 |
| CGCGTGGA | ? | 15 |
| GGATGAGA | ? | 1 |

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 \Rightarrow


Which DNA sequences belong to the same species?

This can be described by a **Poisson mixture model**.

- 1 The Poisson Distribution
- 2 Mixture Models
- 3 Expectation-Maximization
- 4 Wrap-up

Multiple Bernoulli trials

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| C_1 | C_2 | $P(C_1, C_2)$ |
|-------|-------|-----------------------------------|
| H | H | $\theta \cdot \theta$ |
| H | T | $\theta \cdot (1 - \theta)$ |
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So, $P(\text{exactly one H}) = 2 \cdot \theta \cdot (1 - \theta)$.

In general, $P(\text{exactly } m \text{ successes in } n \text{ trials}) = \binom{n}{m} \cdot \theta^m \cdot (1 - \theta)^{n-m}$.

Take it, to the limit, one more time

What if we have an infinite number of trials and expect to see λ successes?

$$\lim_{n \rightarrow \infty} P(\text{exactly } m \text{ successes in } n \text{ trials}) = \frac{\lambda^m}{m!} \exp\{-\lambda\}$$

This is called the **Poisson distribution**.

We will write $g(m : \lambda)$ to mean $P(\text{exactly } m \text{ successes given } \lambda)$.

(See the videos for a detailed derivation.)

Mixtures of distributions

Suppose we have K Poisson distributions (**components**) with parameters $\lambda_1 \dots \lambda_K$ **mixed** together with proportions $p_1 \dots p_K$.

We often write $P = \{p_1 \dots p_K\}$ and $\theta = \{\lambda_1 \dots \lambda_K, P\}$.

```

procedure GENERATEDATASET(Poisson parameters  $\lambda_1 \dots \lambda_k$ , mixing proportions
 $p_1 \dots p_k$ , samples  $N$ )
   $\mathcal{D} \leftarrow \emptyset$ 
  for  $l = 1$  to  $N$  do
    component  $z_l \leftarrow \text{sample}(\text{Mult}(p_1 \dots p_K))$ 
    observation  $D_l \leftarrow \text{sample}(\text{Poisson}(\lambda_{z_l}))$ 
     $\mathcal{D} \leftarrow \mathcal{D} \cup D_l$ 
  end for
  return  $\mathcal{D}$ 
end procedure

```

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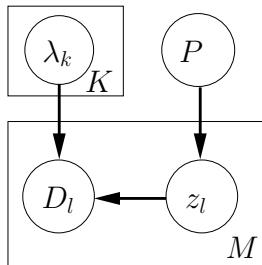


Figure: Generative model for a Poisson mixture model (PMM)

Likelihood of data

We can write the (log) probability of any mixture model as follows.

$$P(\mathcal{D} : \theta) = \sum_k^K p_k g(\mathcal{D} : \lambda_k)$$

$$P(\mathcal{D} : \theta) = \prod_l^N \sum_k^K p_k g(D_l : \lambda_k)$$

$$\ell(\mathcal{D} : \theta) = \log \prod_l^N \sum_k^K p_k g(D_l : \lambda_k)$$

$$\ell(\mathcal{D} : \theta) = \sum_l^N \log \sum_k^K p_k g(D_l : \lambda_k)$$

The learning problem can be formulated as follows.

$$\theta^* = \arg \max_{\theta} \ell(\mathcal{D} : \theta)$$

Membership probabilities

Notation

$q(k, l) := p_k g(D_l : \lambda_k)$ joint probability of D_l and component k

$P(k|l) := P(z_l = k | D_l)$ conditional probability of component k given D_l

The probability that D_l came from component k is expressed as follows.

$$P(k|l) = \frac{q(k, l)}{\sum_m^K q(m, l)}$$

Also, we know each observation came from *some* component.

$$\sum_k P(k|l) = 1$$

Jensen's Inequality

Recall the likelihood of the mixture model.

$$\ell(\mathcal{D} : \theta) = \sum_l^N \log \sum_k^K q(k, l)$$

Jensen's inequality shows the following.

$$\log \sum_k^K \pi_k \alpha_k \geq \sum_k^K \pi_k \log \alpha_k \quad \text{when } \pi \text{ is a distribution}$$

We can make this work for any values.

$$\log \sum_k^K c_k = \log \sum_k^K c_k \frac{\pi_k}{\pi_k} = \log \sum_k^K \pi_k \frac{c_k}{\pi_k} \geq \sum_k^K \pi_k \log \frac{c_k}{\pi_k}$$

Expectation-Maximization (EM)

Our learning problem is formulated as follows.

$$\theta^* = \arg \max_{\theta} \ell(\mathcal{D} : \theta)$$

EM begins with a (bad) set of estimates for θ .

- 1 Use Jensen's inequality to estimate a bound b on ℓ called the **expectation** of ℓ
- 2 Find values of θ which **maximize** b

EM is guaranteed to find θ s which do not decrease b .

Expectation and the Q function

Recall the definition of ℓ and Jensen's inequality.

$$\begin{aligned}\ell(\mathcal{D} : \theta) &= \sum_l^N \log \sum_k^K q(k, l) \\ &\geq \sum_l^N \sum_k^K P(k|l) \log \frac{q(k, l)}{P(k|l)}\end{aligned}$$

This gives the **expectation** of ℓ with our current parameters θ .

Based on this equation, we define $Q(\theta)$ which we want to maximize.

$$Q(\theta) = \sum_l^N \sum_k^K P(k|n) \log q(k, l)$$

(See the handout for a detailed derivation of Q .)

Maximization and the Q function

We use the following process to maximize Q for a particular parameter θ_i .

- 1 Differentiate Q w.r.t θ_i
- 2 Set the derivative equal to 0
- 3 Solve for θ_i

(See the handout for detailed derivations.)

$$\lambda_k = \frac{\sum_l^N P(k|l)D_l}{Z(k)}$$
$$p_k = \frac{Z(k)}{N}$$

The EM algorithm for PMMs

procedure PMMEM(data \mathcal{D} , initial $p_1 \dots p_K, \lambda_1 \dots \lambda_K$, convergence criteria \mathcal{C})

while \mathcal{C} has not been met **do**

$$q(k, l) \leftarrow p_k \cdot g(D_l, \lambda_k)$$

$$P(k|l) \leftarrow \frac{q(k,l)}{\sum_m^K q(m,l)}$$

$$\lambda_k \leftarrow \frac{\sum_l^N P(k|l) D_l}{Z(k)}$$

$$p_k \leftarrow \frac{Z(k)}{N}$$

end while

end procedure

▷ Update the expectations

▷ Maximize the parameters

Grouping the DNA sequences into clusters

After running EM, we have several useful pieces of information about our metagenomics sample.

- $P(k|I)$. The distribution over species for each sequence.
- p_k . The relative genome sizes of the species.
- λ_k . The abundance of the species.

Other questions...

- Do we really know how many species there are?
- Can we differentiate species with similar abundances?
- How do we pick “good” initial parameters?
- When have we converged?

More on EM

EM is a general framework that is useful whenever data is missing.

- If used to estimate class probabilities in naive Bayes models, it is called Bayesian clustering
- If used in HMMs, it is called the Baum-Welch algorithm
- Can be used in general Bayesian networks to calculate parameters when some data is missing
- If used with structure learning algorithms, it is called Structural EM
- Many, many others...

We maximize likelihood with EM. What if we want MAP parameters?

Recap

During this part of the course, we have discussed:

- Mixture models as a probabilistic clustering method
- Expectation-maximization as a framework for estimating parameters when variables are hidden

Next in probabilistic models

We will see a Bayesian version of EM.

- Estimating parameters in topic models