

On the Identifiability and Estimation of Causal Effects

Machine Learning Coffee Seminar

Antti Hyttinen, University Researcher

Jussi Viinikka, Doctoral Student

Sums of Products research group
HIIT, Department of Computer Science, University of Helsinki

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- 1 Background: Causal Effects and Non-parametric Identifiability
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- 3 Background: Linear Causal Effect Estimation from Data
- 4 Bayesian Posteriors for Linear Causal Effects
- 5 Gadget: Scalable MCMC Sampling of DAGs
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Section 2:

S. Tikka, A. Hyttinen, and J. Karvanen.

Identifying causal effects via context-specific independence relations. In NeurIPS, 2019.

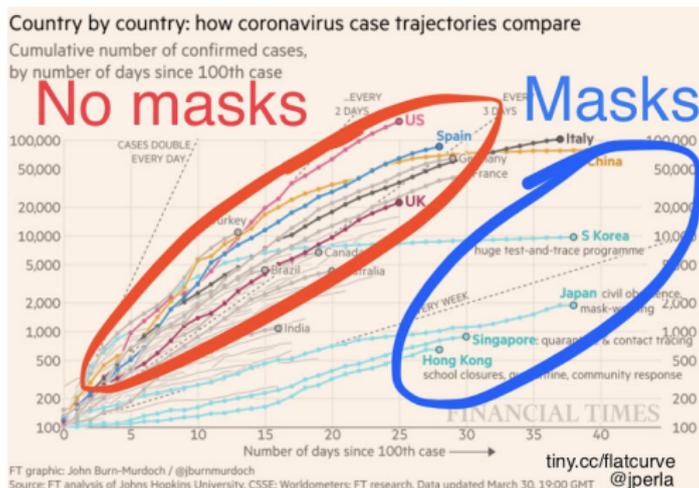
Sections 4 & 5:

J. Viinikka, A. Hyttinen, J. Pensar, and M. Koivisto.

Towards scalable Bayesian learning of causal DAGs.
In NeurIPS, 2020.

Background: Causal Effects and Non-parametric Identifiability

The Need for Quantifying Causal Effects



- Correlation does not imply causation.
- How large is the causal effect?
- Lack of evidence vs. evidence for an insignificant effect.

Causal Effects

- **Causal effects** are probability distributions e.g:

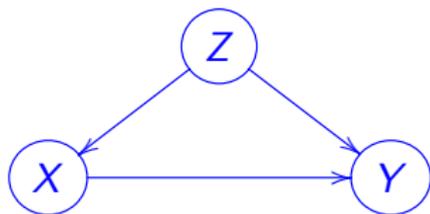
$$P(Y|\text{do}(X)),$$

where $\text{do}(X)$ intervenes at X and sets it to e.g. x .

- E.g. $P(\text{cancer}|\text{do}(\text{smoke}))$, $P(\text{Infection}|\text{do}(\text{wear a mask}))$.
- Randomized controlled trials (RCTs) are a direct way to obtain $P(Y|\text{do}(X))$.
 - Infeasible? Unethical? Expensive? Sample size? Population?

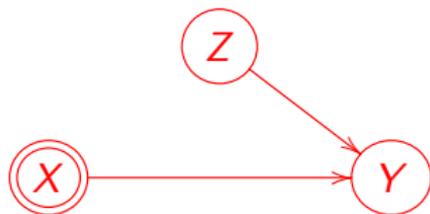
Causal Bayesian Networks

Passive observation



$$P(X, Y, Z) = P(Z)P(X|Z)P(Y|Z, X)$$

do(X)



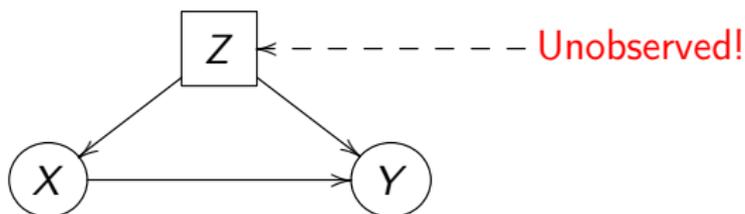
$$P(Y, Z|do(X)) = P(Z)P(Y|Z, X)$$

- Edges in the DAG denote direct causal relationship.
- CPDs define **stochastically** how each variable gets its value based on its direct causes.
- Dependence corresponds to reachability in the graph.
- **Intervention corresponds to edge breaking, or dropping from the factorization.**
- Generally $P(Y|do(X)) \neq P(Y|X)$ (doing vs. seeing).

Determining Causal Effects?

- ① What data do you have?
 - interventions (RCTs)?
 - missing data?
- ② What background knowledge do you have?
 - causal graph?
- ③ Which assumptions you are willing to make?
 - acyclicity?
 - Causally sufficient or latent confounders?
 - parametric restrictions, e.g., linearity?
 - selection bias?
- ④ Which output do you want?
 - Identifiability?
 - estimation?
 - bounds?
 - Average causal effect?

Causal Effect Identifiability [Pearl, 2000]



Problem (Causal Effect Identifiability)

*Input: A DAG over V ,
passively observed $P(W)$ for $W \subseteq V$,
a query $P(Y|do(X))$.*

*Task: Output a formula for $P(Y|do(X))$ over $P(W)$,
or decide that it is non-identifiable.*

- Can the effect $P(Y|do(X))$ be uniquely computed from $P(\cdot)$?
- Or, are there two different parameterizations that yield same $P(\cdot)$ but different $P(Y|do(X))$?
- Aim for a general and complete theory!

Do-Calculus [Pearl, 1995]

Rule 1 (Insertion/deletion of observations):

$$P(Y|do(X), Z, W) = P(Y|do(X), W)$$

if $Y \perp\!\!\!\perp Z | X, W$ in $G_{\overline{X}}$
(edges into X removed)

Rule 2 (Action/observation exchange):

$$P(Y|do(X), do(Z), W) = P(Y|do(X), Z, W)$$

if $Y \perp\!\!\!\perp Z | X, W$ in $G_{\overline{X}, \underline{Z}}$
(edges into X removed,
edges out of Z removed)

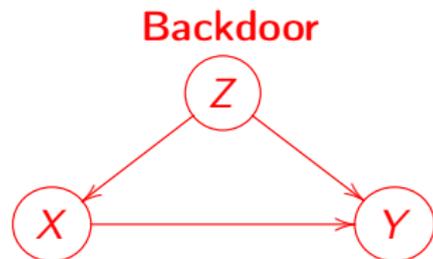
Rule 3 (Insertion/deletion of actions):

$$P(Y|do(X), do(Z), W) = P(Y|do(X), W)$$

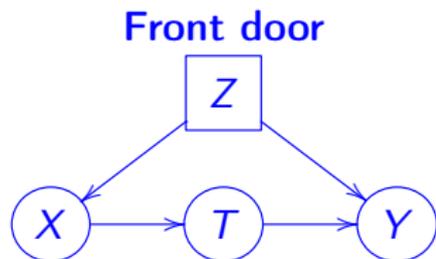
if $Y \perp\!\!\!\perp Z | Z, W$ in $G_{\overline{X}, \overline{Z(W)}}$
(edges into X removed,
and in it edges into Z s that
are not ancestors of W removed)

- Together with probability calculus!

Identifiability and the ID-algorithm [Tian and Pearl, 2002, Shpitser and Pearl, 2006a]



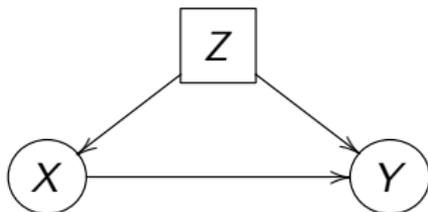
$$P(Y|do(X)) = \\ = \sum_Z P(Y|Z, X)P(Z)$$



$$P(Y|do(X)) = \\ \sum_T P(T|X) \sum_{X'} P(C|T, X')P(X')$$

- The ID-algorithm can find the expressions in polynomial time.
- Use probabilistic modelling (e.g. BN) to calculate the terms.

Non-identifiability and Completeness



- Impossible to untangle the dependence through $X \rightarrow Y$ from the dependence through $X \leftarrow Z \rightarrow Y$.
- A graphical object called **hedge** witnesses **non-identifiability**.

#1	#2
$P(Z = 1) = 0.5$	$P(Z = 1) = 0.5$
$P(X = 1 Z) = 0.5$	$X Z = Z$
$Y X, Z = X$	$Y X, Z = Z$
$P(X, Y = X) = 0.5$	$P(X, Y = X) = 0.5$
$P(Y = 1 do(X = 1)) = 1$	$P(Y = 1 do(X = 1)) = 0.5$

- ID and do-calculus are **complete**. [Shpitser and Pearl, 2006a, Huang and Valtorta, 2006]

State of the Art in Non-parametric Identifiability

	Problem (Reference)	Target	Input (assumptions)	Missing data pattern	Method (complete)
1	CE identifiability [Shpitser and Pearl, 2006a]	$P(Y do(X))$	$P(W)$	None	ID (Yes)
2	CE identifiability [Shpitser and Pearl, 2006b]	$P(Y do(X), Z)$	$P(W)$	None	IDC (Yes)
3	z-identifiability [Bareinboim and Pearl, 2012]	$P(Y do(X), Z)$	$P(W), P(W \setminus B do(B))$ (NE, ED)	None	zID (Yes)
4	g-identifiability [Lee et al., 2019]	$P(Y do(X))$	$\{P(W \setminus B_i do(B_i))\}$ (ED)	None	gID (Yes)
5	Surrogate outcome [Tikka and Karvanen, 2019]	$P(Y do(X), Z)$	$\{P(A_i do(B_i), C_i)\}$ (NE, SO)	None	TRSO (No)
6	mz-transportability [Bareinboim and Pearl, 2014]	$P(Y do(X), Z)$	$\{P(W \setminus (B_i \cup T_i) do(B_i), T_i)\}$ (NEDD, ED)	None	TR ^{mz} (Yes)
7	Selection bias [Bareinboim and Tian, 2015]	$P(Y do(X), Z)$	$P(W \setminus S S)$	Selection	RC (Unknown)
8	Gen. identifiability [Tikka et al., 2020]	$P(Y do(X), Z)$	$\{P(A_i do(B_i), C_i)\}$	None	do-search (Unknown)
9	Missing data [Mohan et al., 2013]	$P(W)$	$P(W^*)$	Restricted	- (Yes)
10	Missing data [Bhattacharya et al., 2019]	$P(W)$	$P(W^*)$	Arbitrary	- (Unknown)
11	Gen. identifiability [Tikka et al., 2020]	$P(Y do(X), Z)$	$\{P(A_i^* do(B_i), C_i^*)\}$	Arbitrary	do-search (No)

Identifying causal effects via CSI relations

Context-specific Independence [Boutilier et al., 1996]

$$X \perp\!\!\!\perp Y|Z = 0$$

i.e. $P(X|Y, Z = 0) = P(X|Z = 0)$

but $X \not\perp\!\!\!\perp Y|Z = 1$ (possibly)

- A very natural independence restriction, for example:

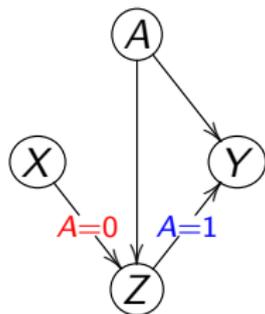
$$\text{INCOME} \perp\!\!\!\perp \text{WEATHER}|\text{JOB} = \text{clerk}$$

$$\text{INCOME} \not\perp\!\!\!\perp \text{WEATHER}|\text{JOB} = \text{farmer}$$

- CSIs have been extensively exploited in BN inference, but only recently been used to make novel causal inferences. [Hyttinen et al., 2018, Mooij et al., 2020]

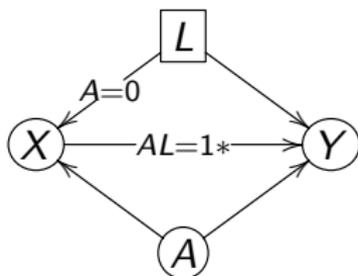
Labeled DAGs [Pensar, Nyman, Koski, and Corander, 2015]

$P(Z A, X)$	$Z = 0$	$Z = 1$
$AX = 00$	0.1	0.9
$AX = 01$	0.1	0.9
$AX = 10$	0.5	0.5
$AX = 11$	0.6	0.4



- A label on an edge encodes contexts where the edge is absent.
- Any assignment in a label denotes a local CSI:
e.g. $X \perp\!\!\!\perp Z|A = 0$.
- Labels allow for representation, theory on equivalence classes, and separation criteria.

Causal Effect Identifiability via CSIs [Tikka, Hyttinen, and Karvanen, 2019]



Problem (Causal Effect Identifiability via CSIs)

Input: An **LDAG** over V ,
passively observed $P(W)$ for $W \subseteq V$,
a query $P(Y|do(X))$.

Task: Output a formula for $P(Y|do(X))$ over $P(W)$,
or decide that it is non-identifiable.

CSI-do-calculus [Tikka, Hyttinen, and Karvanen, 2019]

Rule 1 (Insertion/Deletion of observations):

$$P(Y_1, y_2 | Z_1, z_2, X_1, x_2) = P(Y_1, y_2 | X_1, x_2) \text{ if } Y_1, Y_2 \perp\!\!\!\perp Z_1, Z_2 | X_1, x_2$$

Rule 2 (Marginalization/Sum-rule): $P(Y_1, y_2 | X_1, x_2) = \sum_Z P(Y_1, y_2, Z | X_1, x_2)$

Rule 3 (Conditioning): $P(Y_1 | Z_1, z_2, X_1, x_2) = \frac{P(Y_1, Z_1, z_2 | X_1, x_2)}{\sum_{Y_1} P(Y_1, Z_1, z_2 | X_1, x_2)}$

Rule 4 (Product-rule): $P(Y_1, y_2, Z_1, z_2 | X_1, x_2) = P(Y_1, y_2 | Z_1, z_2, X_1, x_2) P(Z_1, z_2 | X_1, x_2)$

Rule 5 (General-by-case): $P(Y_1, y_2, 1 - z | X_1, x_2) = P(Y_1, y_2 | X_1, x_2) - P(Y_1, y_2, z | X_1, x_2)$

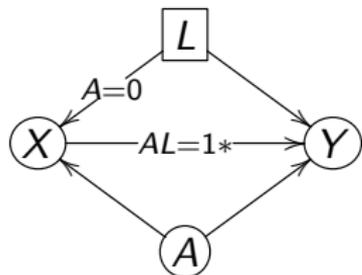
Rule 6 (Case-by-case): $P(Y_1, y_2, Z | X_1, x_2) = \begin{cases} P(Y_1, y_2, Z = 0 | X_1, x_2) \\ P(Y_1, y_2, Z = 1 | X_1, x_2) \end{cases}$

Rule 7 (Case-by-general (a)): $P(Y_1, y_2, z | X_1, x_2) = P(Y_1, y_2, Z | X_1, x_2) \Big|_{Z=z}$

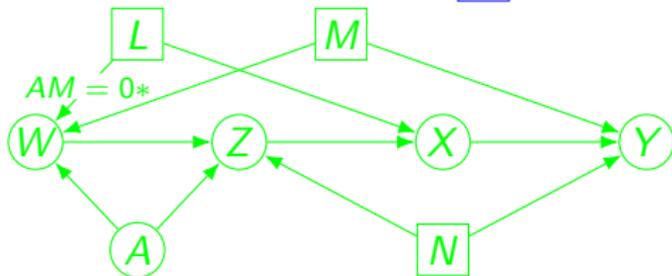
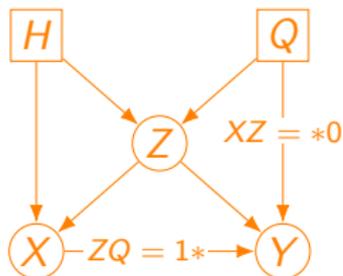
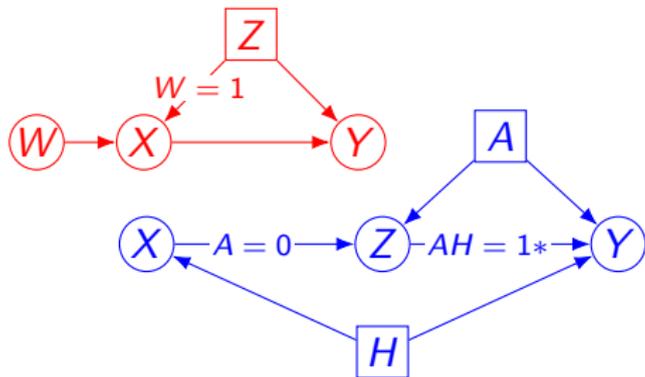
Rule 8 (Case-by-general (b)): $P(Y_1, y_2 | X_1, x_2, z) = P(Y_1, y_2 | X_1, x_2, Z) \Big|_{Z=z}$

- These rules subsume do-calculus, do-operator not needed.
- Deciding non-identifiability is NP-hard: we need case by case reasoning.

Extended Identifiability for $P(Y|do(X))$ via CSIs



$$P(Y|A=0, X)P(A=0) \\ + P(Y|A=1)P(A=1)$$



→ Few CSIs may be sufficient to turn a previously non-identifiable instance into identifiable.

Background: Linear Causal Effect Estimation from Data

Linear Causal Effect Estimation from Data

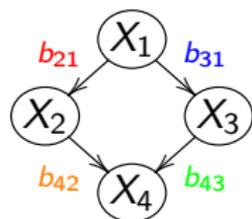
Problem (Linear Causal Effect Estimation from Data)

*Input: Passively observed **causally sufficient** data D .*

*Task: Estimate the **linear** causal effect*

$$\pi_{ji} = \frac{\partial}{\partial x_i} E(X_j | do(X_i = x_i))$$

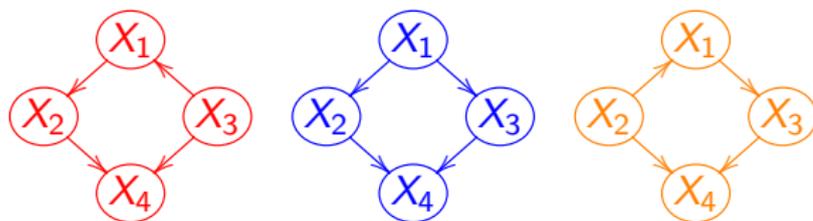
- Sum-product of edge coefficients on directed paths.



$$\begin{aligned}\pi_{41} &= b_{42} \cdot b_{21} + b_{43} \cdot b_{31} \\ &= -1.03 \cdot .78 + .74 \cdot .60 = -.36\end{aligned}$$

IDA [Maathuis, Kalisch, and Bühlmann, 2009]

- Gaussian DAGs only identifiable up to the Markov eq. class.



- For a fixed DAG, the causal effect is identifiable via linear regression over the cause X_1 and its parents (backdoor adj.):

$$X_4 \sim X_1 + X_3, \quad X_4 \sim X_1, \quad X_4 \sim X_1 + X_2$$

- Output is a set of possible causal effects:

$$\pi_{41} \in \{ -.74, . - .29, .49 \}$$

- Need only the possible parent sets of the cause X_1 via e.g. PC.

Bayesian Posteriors for Linear Causal Effects

BIDA [Pensar, Talvitie, Hyttinen, and Koivisto, 2020, AAAI]

Posterior of π_{ji}

Bayesian lin. regression of X_j over $X_i, pa(X_i)$.

computed **exactly** via dynamic prog.

$$f(\pi_{ji}|\mathbf{D}) = \sum_{pa(X_i)} f(\pi_{ji}|\mathbf{D}, pa(X_i))P(pa(X_i)|\mathbf{D})$$

Bayesian model averaging over DAGs

Beeps: Bayesian Effect Estimation by Posterior Sampling

[Viinikka, Hyttinen, Pensar, and Koivisto, 2020, NeurIPS]

Posterior of π_{ji}

Bayesian model averaging over DAGs

Sample G using MCMC

$$f(\pi_{ji} | \mathbf{D}) = \sum_{G, \mathbf{B}} f(\pi_{ji} | \mathbf{B}) f(\mathbf{B} | \mathbf{D}, G) P(G | \mathbf{D})$$

Sample coefficients \mathbf{B} (BGe prior)

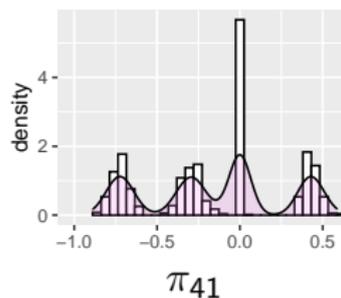
$$\pi_{ji} = ((\mathbf{I} - \mathbf{B})^{-1})[j, i]$$

IDA example continued...

IDA

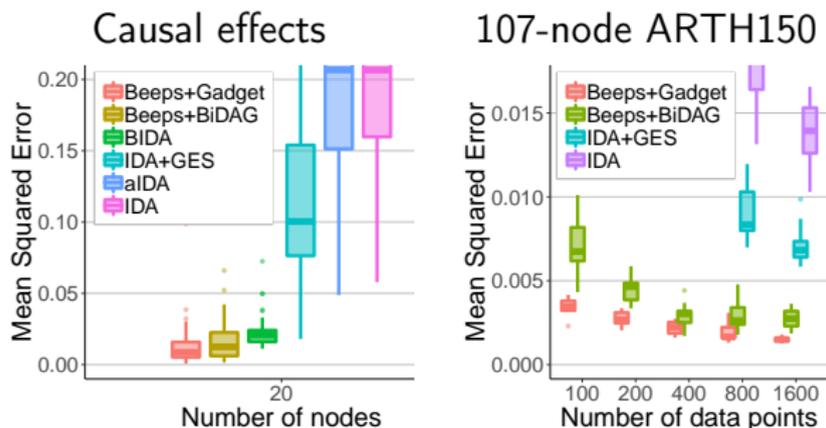
$$\pi_{41} \in \{-.74, -.29, .49\}$$

Beeps



- Uncertainty in the estimates due to low sample size.
 - Uncertainty over the causal structure.
- Beeps catches both.

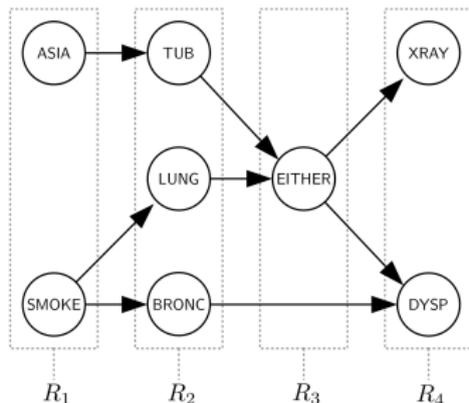
Simulation Results



- Better accuracy than IDA-based methods.
- We can scale up to over 100 nodes and our MCMC outperforms BiDAG.
- Concurrent similar suggestions by Kuipers et al. [2019] and Castelletti and Consonni [2020].

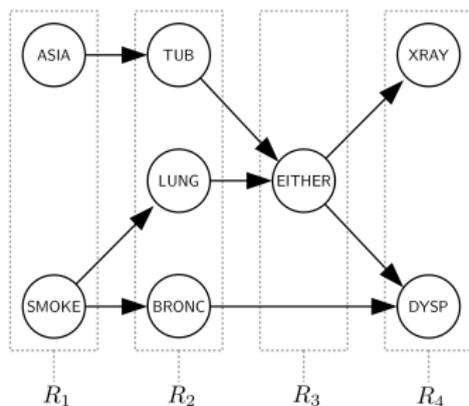
Gadget: Scalable MCMC Sampling of DAGs

Background – Root-partitions (Kuipers & Moffa, 2017)



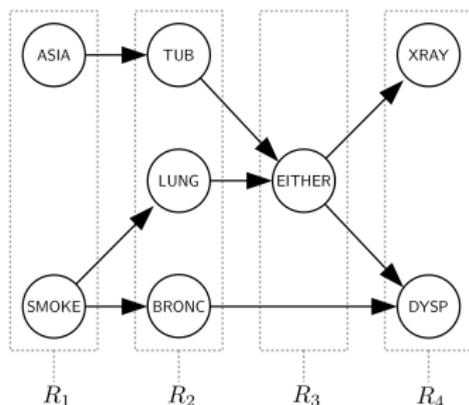
- Root-partition R : ordered set partition of DAG nodes V
- Iterate: root-nodes to the next part, remove them from the DAG
- For $v \in R_i$ parents valid if $pa(v) \subseteq \underbrace{R_{1,i-1}}_U$, $pa(v) \cap \underbrace{R_{i-1}}_T \neq \emptyset$
- $\pi(R) = \prod_{t=1}^k \prod_{i \in R_t} \tau_i(R_{1,t-1}, R_{t-1})$, with $\tau_i(U, T) := \sum_{S \subseteq U: S \cap T \neq \emptyset} \pi_i(S)$

Background – K candidate parents (Kuipers et al., 2020)



- For each node $i \in V$ select K candidate parents C_i
- Now $\tau_i(U, T) = \tau_i(U \cap C_i, T \cap C_i)$, $O(3^K)$ space per node

Preprocessing – Scoring nodes by subtraction

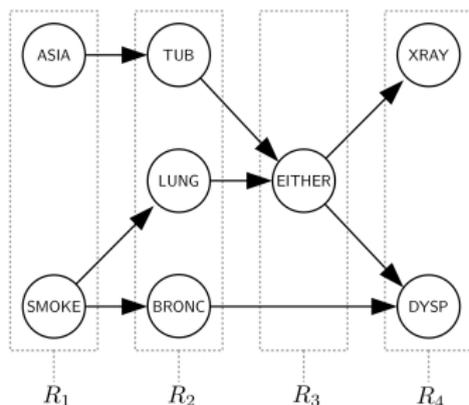


- For each node $i \in V$ select K candidate parents C_i
- Now $\tau_i(U, T) = \tau_i(U \cap C_i, T \cap C_i)$, $O(3^K)$ space per node
- For any $i \in V$ and $J \subseteq V \setminus \{i\}$, let

$$\tau_i(J) := \sum_{S \subseteq J \cap C_i} \pi_i(S)$$

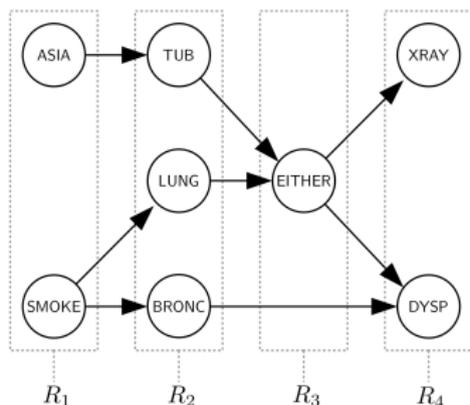
- Now $\tau_i(U, T) = \tau_i(U) - \tau_i(U \setminus T)$, $O(2^K)$ space per node

Preprocessing – Parents outside of the candidates



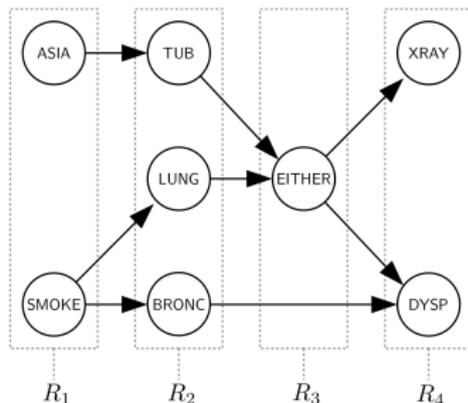
- Need to allow small number of parents outside of candidates too
 - i. C_i may not be optimal or large enough for all $i \in V$
 - ii. Posterior landscape may contain large zero-probability regions making transition between root-partitions inefficient for the MCMC
- We allow parent sets of maximum indegree d not contained in C_i
- Scores sorted, only need to accumulate certain amount to reach acceptable error

Markov chain – Metropolis Coupled Markov Chain Monte Carlo



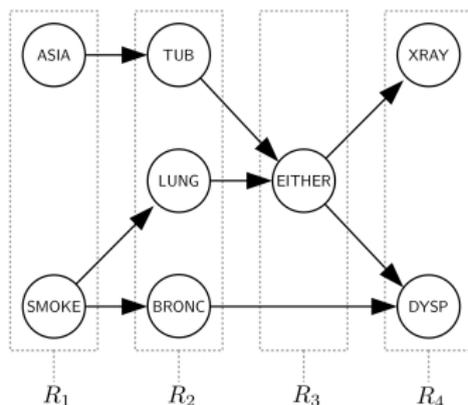
- Moves in root-partition space:
 - split R_i
 - merge R_i and R_{i+1}
 - swap nodes between R_i and R_j
- M parallel “heated” chains with the stationary distribution of k th chain proportional to $\pi^{k/M}$, i.e. Metropolis coupling
- On every other step a randomly chosen chain and neighbour swap states with certain probability

Post processing – Sampling DAGs



- DAG is sampled for each root-partition stored during Markov chain simulation
- Datastructure allowing fast parent set sampling after time and space invested in precomputing
- Sampling proceeds one node at a time for all DAGs generated

Gadget – Summary



1. Preprocessing

- Candidate parent selection
- Precomputing data structures for scoring root-partitions

2. Markov chain simulation

3. Postprocessing

- Generating DAGs from root-partitions

Conclusion

Research directions

- Unknown graph and latent variables. [Hyttinen et al., 2015, Malinsky and Spirtes, 2017, Jaber et al., 2019]
- Soft interventions. [Correa and Bareinboim, 2020]
- Cyclic causal graphs. [Forré and Mooij, 2019]
- Intervals and bounds. [Malinsky and Spirtes, 2017, Peters et al., 2016]
- Path specific effects. [Malinsky et al., 2019]
- Linear identifiability. [Kumor et al., 2020]
- Counterfactuals. [Kusner et al., 2017, Shpitser and Pearl, 2008]

Conclusion

- Need causal effects from the data and knowledge we have.
- CSIs allow for identifiability beyond do-calculus.
- Bayesian posteriors for linear causal effects are more accurate and characterize the remaining uncertainty.
- State of the art MCMC posterior sampling for DAGs.
- Future work:
 - Completeness of CE identification via CSIs?
 - Relax the assumptions for Bayesian posteriors?
 - How to select candidate parents? How to scale up further?

Collaborators: Mikko Koivisto (UH), Johan Pensar (University of Oslo), Santtu Tikka (University of Jyväskylä), Juha Karvanen (University of Jyväskylä), Topi Talvitie (UH)

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