## 582631 Introduction to Machine Learning, Fall 2016 Exercise set II

## Model solutions

1. 

(a) Because the points are uniformly distributed, the expected fraction of the points residing on a subinterval is the length of the subinterval compared to the length of the whole interval. If we denote the length of the interval by $\alpha$, the expected number of points falling within the interval, $E(\alpha)$, is given by

$$
E(\alpha)=\frac{0.1}{1}=0.1 .
$$

(b) In two dimensions, we have likewise that the expected number of points is given by the relative area of the smaller square compared to the area of the whole square:

$$
E(\alpha)=\frac{0.1^{2}}{1^{2}}=0.01 .
$$

(c) For general $p \geq 1$, the expected fraction is the relative volume of the smaller hypercube compared to the volume of the whole hypercube:

$$
E(\alpha)=\frac{0.1^{100}}{1^{100}}=10^{-100}
$$

(d) If we assume the observations are uniformly distributed and use aforementioned definition of "nearness", on average only $10^{-p}$ training observations are near the given test observation when the number of features is $p \geq 1$.
In practice, things may turn out not to be as bad as this: observations are usually not uniformly distributed, and sometimes two observations do not have to be near each other in all $p$ dimensions to be considered similar to each other - for example, some of the features may be irrelevant in view of the task at hand.
(e) We can solve the length, $x$, of the side of the hypercube that contains on the average $10 \%$ of the data:

$$
\begin{aligned}
\frac{1}{10} & =E(\alpha)=\frac{x^{p}}{1^{p}} \\
x & =10^{-1 / p} .
\end{aligned}
$$

For $p=1,2$ and 100 , this is:

$$
\begin{aligned}
& x=10^{-1}=0.1 \\
& x=10^{-1 / 2} \approx 0.3162 \\
& x=10^{-1 / 100} \approx 0.9772 .
\end{aligned}
$$

So the hypercube that contains on the average only $10 \%$ of the data has sides of length approximately 0.9772 , so almost 1 ! This demonstrates that data is sparse in highdimensional spaces. In other words, the training observations that are among the $10 \%$ of the training data nearest to the test observation may actually be almost maximally different from the test observation.
2. Define a binary random variable $D$ as follows:

$$
D= \begin{cases}1, & \text { if a company pays a dividend } \\ 0, & \text { otherwise }\end{cases}
$$

The prior probabilities for $D$ are $\operatorname{Pr}(D=0)=\pi_{0}=0.2$, and $\operatorname{Pr}(D=1)=\pi_{1}=0.8$.
The conditional distributions of the percentage profit $X$ given values of $D$ are:

$$
\begin{aligned}
& f_{X \mid D=0}(x)=\mathcal{N}\left(x ; \mu_{0}, \sigma^{2}\right) \\
& f_{X \mid D=1}(x)=\mathcal{N}\left(x ; \mu_{1}, \sigma^{2}\right)
\end{aligned}
$$

where the expected values are $\mu_{0}=0, \mu_{1}=10$, and a common variance is $\sigma^{2}=36$.
We get the probability that a company pays a dividend given that its percentage profit is $X=4$ using the Bayes' theorem ${ }^{11}$

$$
\begin{aligned}
P(D=1 \mid X=4) & =\frac{\pi_{1} f_{X \mid D=1}(4)}{\pi_{1} f_{X \mid D=1}(4)+\pi_{0} f_{X \mid D=0}(4)} \\
& =\frac{0.8 \cdot \frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left(-\frac{\left(4-\mu_{1}\right)^{2}}{2 \sigma^{2}}\right)}{0.8 \cdot \frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left(-\frac{\left(4-\mu_{1}\right)^{2}}{2 \sigma^{2}}\right)+0.2 \cdot \frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left(-\frac{\left(4-\mu_{0}\right)^{2}}{2 \sigma^{2}}\right)} \\
& =\frac{0.8 e^{-\frac{1}{2}}}{0.8 e^{-\frac{1}{2}}+0.2 e^{-\frac{2}{9}}} \approx 0.752
\end{aligned}
$$

Thus, the probability that the company pays a dividend given that its percentage profit is 4.0 is about $75 \%$.

Note that this is just slightly lower than the prior probability, $\operatorname{Pr}(D=1)$, which was $80 \%$. This is consistent with the reasoning that the average percentage profit of companies that pay dividends, $\mu_{1}=10.0$, is a little further away from the observed profit, $x=4.0$, than the average profit of companies that don't pay dividends, $\mu_{0}=0.0$. So while the observed profit is slightly more typical to non-dividend-paying companies, the difference is so small that the effect on the conditional probability is less than 5 percentage points.

We are of course curious to know what the posterior probability would have looked like if the profit percentage, $x$, had been something different. To look into this, let's write a little R script that computes the posterior $P(D=1 \mid X=x)$ for different $x$ :

```
po <- function(x) { a=.8*exp(-(x-10)~2/(2*36)); b=.2*exp(-(x-0)~2/(2*36)); a/(a+b) }
x = .5*(-30:50) # a grid of points between -15 and 25
plot(x, po(x)) # basic plot showing the posterior as a function of x
plot(x, po(x), t='h', lwd=6, col='orange', lend=2) # fancy plot showing the same
points(4, po(4),t='h',lwd=6,col='black',lend=2) # emhasize x=4 in black
```

(plot shown on the next page...)

[^0]

A fancy plot of the posterior with the point $x=4$ highlighted in black.
As you can see, the model predicts that the chances that a company pays dividends decrease as the profit $x$ decreases, and vice versa. What we have managed to implement here is actually one-dimensional Linear Discriminant Analysis (LDA)!

In next week's exercises, we'll do something similar in two-dimensions.


[^0]:    ${ }^{1}$ Bayes' theorem applies also for the joint distribution of the discrete and continuous random variable. For a continuous variable probabilities in the formula are replaced by densities, and conditional probabilities by conditional densities.

